# **2D COMPUTER GRAPHICS**

Diego Nehab

Summer 2020

IMPA

## ANTI-ALIASING AND TEXTURE MAPPING

Let f be a function and  $\psi$  an *anti-aliasing filter* 

#### ANTI-ALIASING

Let f be a function and  $\psi$  an anti-aliasing filter

Value of pixel  $p_i$  is given by

$$p_i = (f * \psi)(i) = \int_{-\infty}^{\infty} f(t) \psi(i-t) dt$$

Let f be a function and  $\psi$  an anti-aliasing filter

Value of pixel  $p_i$  is given by

$$p_i = (f * \psi)(i) = \int_{-\infty}^{\infty} f(t) \psi(i - t) dt$$

How to compute the integral when f is a vector graphics illustration?

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

Possible to clip edges, not the shapes

- + general piecewise polynomial filters [Duff, 1989]
- + curved edges [Manson and Schaefer, 2013]

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

Possible to clip edges, not the shapes

- + general piecewise polynomial filters [Duff, 1989]
- + curved edges [Manson and Schaefer, 2013]

What about polygons with self-intersections?

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

Possible to clip edges, not the shapes

- + general piecewise polynomial filters [Duff, 1989]
- + curved edges [Manson and Schaefer, 2013]

What about polygons with self-intersections?

What about spatially varying colors?

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

Possible to clip edges, not the shapes

- + general piecewise polynomial filters [Duff, 1989]
- + curved edges [Manson and Schaefer, 2013]

What about polygons with self-intersections?

What about spatially varying colors?

What about multiple opaque layers?

- Clip polygon against the **box** centered at each pixel
- Compute weighted area using on Green's theorem from Calculus

Possible to clip edges, not the shapes

- + general piecewise polynomial filters [Duff, 1989]
- + curved edges [Manson and Schaefer, 2013]

What about polygons with self-intersections?

What about spatially varying colors?

What about multiple opaque layers?

What about transparency?

Assume path  $P_i$  with constant color  $f_i$ ,  $\alpha_{f_i}$ 

Assume path  $P_i$  with constant color  $f_i$ ,  $\alpha_{f_i}$ 

Assume blending over the background  $b_i, \alpha_{b_i}$ 

Assume path  $P_i$  with constant color  $f_i$ ,  $\alpha_{f_i}$ 

Assume blending over the background  $b_i, \alpha_{b_i}$ 

Assume anti-aliasing filter  $\psi$  with support  $\Omega$ 

#### POPULAR HACK

Assume path  $P_i$  with constant color  $f_i$ ,  $\alpha_{f_i}$ 

Assume blending over the background  $b_i, \alpha_{b_i}$ 

Assume anti-aliasing filter  $\psi$  with support  $\Omega$ 

Define the coverage o of  $P_i$  at pixel p

$$o = \int_{\Omega} [u - p \in P_i] \psi(u) \, du$$

#### POPULAR HACK

Assume path  $P_i$  with constant color  $f_i$ ,  $\alpha_{f_i}$ 

- Assume blending over the background  $b_i, \alpha_{b_i}$
- Assume anti-aliasing filter  $\psi$  with support  $\Omega$

Define the coverage o of  $P_i$  at pixel p

$$o=\int_{\Omega}[u-p\in P_i]\,\psi(u)\,du$$

The new background  $b_{i+1}, \alpha_{i+1}$  is

$$b_{i+1}, \alpha_{i+1} = f_i, (\alpha_i \cdot o) \oplus b_i, \alpha_i$$

#### **PROBLEMS WITH HACK**

Visible seams at perfectly abutting layers, weird halos



This is called the correlated mattes problem

## **PROBLEMS WITH HACK**

Visible seams at perfectly abutting layers, weird halos



This is called the *correlated mattes* problem

It also either blends in linear, or antialiases in gamma





Notice the change in thickness.

Notice the change in thickness.

## PROBLEMS WITH HACK

Visible seams at perfectly abutting layers, weird halos



This is called the *correlated mattes* problem

It also either blends in linear, or antialiases in gamma





Notice the change in thickness.

Notice the change in thickness.

Must blend in gamma and antialias in linear [Nehab and Hoppe, 2008]  $b_{i+1}, \beta i + 1 = \gamma \left( \gamma^{-1}(f_i, \alpha_i \oplus b_i, \beta_i) \cdot o + \gamma^{-1}(b_i, \beta_i) \cdot (1 - o) \right)$ 

## **PROBABILITY IN 2 SLIDES**

A random variable X is a function that maps outcomes to numbers

 $F_X(a) = P[X \leq a]$ 

$$F_X(a) = P[X \leq a]$$

i.e., it measures the probability that the numerical value is at most a.

$$F_X(a) = P[X \leq a]$$

i.e., it measures the probability that the numerical value is at most a.

The associated probability density function  $f_X$  is such that

$$F_X(a) = \int_{-\infty}^a f_X(t) \, dt$$

$$F_X(a) = P[X \leq a]$$

i.e., it measures the probability that the numerical value is at most a.

The associated probability density function  $f_X$  is such that

$$F_X(a) = \int_{-\infty}^a f_X(t) \, dt$$

i.e., its integral is the cumulative distribution function.

$$F_X(a) = P[X \leq a]$$

i.e., it measures the probability that the numerical value is at most *a*.

The associated probability density function  $f_X$  is such that

$$F_X(a) = \int_{-\infty}^a f_X(t) \, dt$$

i.e., its integral is the cumulative distribution function.

The associated expectation E[X] (or mean  $\mu_X$ ) is

$$E[X] = \int_{-\infty}^{\infty} t f_X(t) dt = \mu_X \tag{1}$$

$$F_X(a) = P[X \leq a]$$

i.e., it measures the probability that the numerical value is at most *a*.

The associated probability density function  $f_X$  is such that

$$F_X(a) = \int_{-\infty}^a f_X(t) \, dt$$

i.e., its integral is the cumulative distribution function.

The associated expectation E[X] (or mean  $\mu_X$ ) is

$$E[X] = \int_{-\infty}^{\infty} t f_X(t) dt = \mu_X \tag{1}$$

i.e., the mean value weighted by the probability density function.

## **PROBABILITY IN 2 SLIDES**

The associated variance  $\operatorname{var}(X) = \sigma_X^2$  is  $\operatorname{var}(X) = E[(X - \mu_X)^2] = E[X^2] - E^2[X] = \sigma_X^2$ 

and the standard deviation is  $\sigma_{\chi}$ .

## **PROBABILITY IN 2 SLIDES**

The associated variance  $\operatorname{var}(X) = \sigma_X^2$  is  $\operatorname{var}(X) = E[(X - \mu_X)^2] = E[X^2] - E^2[X] = \sigma_X^2$ 

and the standard deviation is  $\sigma_X$ .

Measure how much the random variable deviates from the mean

The associated variance  $\operatorname{var}(X) = \sigma_X^2$  is  $\operatorname{var}(X) = E[(X - \mu_X)^2] = E[X^2] - E^2[X] = \sigma_X^2$ 

and the standard deviation is  $\sigma_X$ .

Measure how much the random variable deviates from the mean

The sample average is  $\overline{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$ 

The associated variance  $\operatorname{var}(X) = \sigma_X^2$  is  $\operatorname{var}(X) = E[(X - \mu_X)^2] = E[X^2] - E^2[X] = \sigma_X^2$ 

and the standard deviation is  $\sigma_{\chi}$ .

Measure how much the random variable deviates from the mean The sample average is  $\overline{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$ Law of large numbers

 $\overline{X}_n \to \mu_X$  for  $n \to \infty$ 

The associated variance  $\operatorname{var}(X) = \sigma_X^2$  is  $\operatorname{var}(X) = E[(X - \mu_X)^2] = E[X^2] - E^2[X] = \sigma_X^2$ 

and the standard deviation is  $\sigma_X$ .

Measure how much the random variable deviates from the mean The sample average is  $\overline{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$ Law of large numbers

$$\overline{X}_n 
ightarrow \mu_X$$
 for  $n 
ightarrow \infty$ 

Variance of sample average

$$\operatorname{var}(\overline{X}_n) = \operatorname{var}\left(\frac{1}{n}\sum X_i\right) = \frac{1}{n^2}\sum \operatorname{var}(X_i) = \frac{\sigma_X^2}{n}$$

Start by expressing an integral as the expectation of a random variable Estimate expectation by sample mean Start by expressing an integral as the expectation of a random variable

- Estimate expectation by sample mean
- Rely on law of large numbers

Start by expressing an integral as the expectation of a random variable Estimate expectation by sample mean

Rely on law of large numbers

Let X be such that support of  $f_X$  is  $\Omega$ 

$$\int_{\Omega} g(t) dt = \int_{\Omega} \frac{g(t)}{f_{X}(t)} f_{X}(t) dt = E\left[\frac{g(X)}{f_{X}(X)}\right] \approx \frac{1}{n} \sum_{i=1}^{n} \frac{g(X_{i})}{f_{X}(X_{i})}$$

Start by expressing an integral as the expectation of a random variable Estimate expectation by sample mean

Rely on law of large numbers

Let X be such that support of  $f_X$  is  $\Omega$ 

$$\int_{\Omega} g(t) dt = \int_{\Omega} \frac{g(t)}{f_X(t)} f_X(t) dt = E\left[\frac{g(X)}{f_X(X)}\right] \approx \frac{1}{n} \sum_{i=1}^n \frac{g(X_i)}{f_X(X_i)}$$

This is the basis of supersampling

The solution to our anti-aliasing problems

Let  $g: \mathbf{R}^2 \to RGB$  map positions to linear color

Consider an anti-aliasing kernel  $\psi$
Let  $g: \mathbf{R}^2 \to RGB$  map positions to linear color

Consider an anti-aliasing kernel  $\psi$ 

The *linear* color at pixel *p* is

$$c(p) = \int_{\Omega} g(p-q) \, \psi(q) \, dq$$

Let  $g: \mathbb{R}^2 \to RGB$  map positions to linear color Consider an *anti-aliasing kernel*  $\psi$ 

The *linear* color at pixel *p* is

$$c(p) = \int_{\Omega} g(p-q) \psi(q) \, dq$$
$$= E\left[\frac{g(p-X)\psi(X)}{f_X(X)}\right]$$

Let  $g: \mathbf{R}^2 \to RGB$  map positions to linear color Consider an *anti-aliasing kernel*  $\psi$ 

The *linear* color at pixel *p* is

$$F(p) = \int_{\Omega} g(p-q) \psi(q) dq$$
$$= E\left[\frac{g(p-X)\psi(X)}{f_X(X)}\right]$$
$$\approx \frac{1}{n} \sum_{i=1}^{n} \frac{g(p-X_i)\psi(X_i)}{f_X(X_i)}$$

Let  $g: \mathbf{R}^2 \to RGB$  map positions to linear color Consider an *anti-aliasing kernel*  $\psi$ 

С

The *linear* color at pixel p is

$$(p) = \int_{\Omega} g(p-q) \psi(q) dq$$
$$= E\left[\frac{g(p-X)\psi(X)}{f_X(X)}\right]$$
$$\approx \frac{1}{n} \sum_{i=1}^{n} \frac{g(p-X_i)\psi(X_i)}{f_X(X_i)}$$

When  $\psi = \beta^0$  is the box,  $f_X = 1$  with support  $\Omega = [-\frac{1}{2}, \frac{1}{2}]^2$  $c(p) \approx \frac{1}{n} \sum_{i=1}^{n} g(p - X_i)$  Estimator is unbiased if expected value is correct

Estimator is unbiased if expected value is correct

The Monte Carlo estimator is unbiased in this sense

$$c(p) \approx \frac{1}{n} \sum_{i=1}^{n} \frac{g(p - X_i) \psi(X_i)}{f_X(X_i)}$$

Estimator is unbiased if expected value is correct

The Monte Carlo estimator is unbiased in this sense

$$c(p) \approx \frac{1}{n} \sum_{i=1}^{n} \frac{g(p - X_i) \psi(X_i)}{f_X(X_i)}$$

It often makes sense to use a *biased* estimator to reduce variance

$$c(p) \approx \frac{\sum_{i=1}^{n} \frac{g(p - X_i) \psi(X_i)}{f_X(X_i)}}{\sum_{i=1}^{n} \frac{\psi(X_i)}{f_X(X_i)}}$$

What happens if we choose  $f_X(t) \propto g(t)$ ?

# What happens if we choose $f_X(t) \propto g(t)$ ? $\int_{\Omega} g(t) dt = E \left[ \frac{g(X)}{f_X(X)} \right]$

What happens if we choose 
$$f_X(t) \propto g(t)$$
?  
$$\int_{\Omega} g(t) dt = E\left[\frac{g(X)}{f_X(X)}\right] = E[\alpha] = \frac{g(X)}{f(X)}$$

We only need one sample!

What happens if we choose 
$$f_X(t) \propto g(t)$$
?  
$$\int_{\Omega} g(t) dt = E\left[\frac{g(X)}{f_X(X)}\right] = E[\alpha] = \frac{g(X)}{f(X)}$$

We only need one sample!

Unfortunately, we need to normalize g to transform it into a PDF

We only need one sample!

Unfortunately, we need to normalize g to transform it into a PDF

For that, we need to divide it by its integral

This integral is exactly what we are trying to compute!

We only need one sample!

Unfortunately, we need to normalize g to transform it into a PDF

For that, we need to divide it by its integral

This integral is exactly what we are trying to compute!

However, we can often make  $f_X$  almost proportional to g

We only need one sample!

Unfortunately, we need to normalize g to transform it into a PDF

For that, we need to divide it by its integral

This integral is exactly what we are trying to compute!

However, we can often make  $f_X$  almost proportional to g

This is importance sampling

Many different point distributions have  $f_X = 1/A_{\Omega}$  in  $\Omega$ 

Many different point distributions have  $f_X = 1/A_\Omega$  in  $\Omega$ 

Uniform, stratified, low-discrepancy (e.g. Poisson disk, Lloyd relaxation)

Many different point distributions have  $f_X = 1/A_{\Omega}$  in  $\Omega$ Uniform, stratified, low-discrepancy (e.g. Poisson disk, Lloyd relaxation) Variance of  $\overline{X}_n$  is not the same for all of them!

#### **16 SAMPLES**

. . . . . . . . . . . . . . . • . ٠ . . • • . . . • . . . • . . • . • • • • • • • • . • • . . . . . • • • . • • • • • . ٠ . . • ٠ .

Regular

•

.

## 16 SAMPLES

Uniform



## 16 samples

Stratified



.

## 16 samples

a ha a a ha a a ha a h

and the second second

### Blue noise

#### 64 SAMPLES

•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•		•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•
•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•	•

### Regular

Uniform

ŝ •• ۰.

Stratified

2

Blue noise

. 

#### Regular

. .

Uniform

Stratified

Blue noise



#### Regular

Uniform



### Stratified

Blue noise

#### BETTER ANTI-ALIASING KERNELS

1.2 1.2 0.8 0.8 0.6 0.6 0.4 0.4 0.2 0.2 -2 -0.2 -2 -0.2 -42 4 2 4 -420  $-2\pi$ 2π π ±20 0.5 140 . ı. ±60 I.  $-2\pi -\pi 0$ π 2π 180 180




Linear



Gaussian



Keys



Lanczos



#### GENERALIZED SAMPLING



## GENERALIZED SAMPLING



Cardinal cubic B-spline

## GENERALIZED SAMPLING



Cardinal cubic B-spline

Needs sample sharing for variance reduction and speed

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

#### BOX UPSAMPLING



## LINEAR UPSAMPLING



# CARDINAL CUBIC B-SPLINE UPSAMPLING



Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Reconstruction is easy, prefiltering is difficult

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Reconstruction is easy, prefiltering is difficult

Non-uniform resampling

- Reconstruct when locally upsampling
- Prefilter when locally downsampling

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Reconstruction is easy, prefiltering is difficult

Non-uniform resampling

- Reconstruct when locally upsampling
- Prefilter when locally downsampling
- · Jacobian of map from screen to texture coordinates decides

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Reconstruction is easy, prefiltering is difficult

Non-uniform resampling

- Reconstruct when locally upsampling
- Prefilter when locally downsampling
- Jacobian of map from screen to texture coordinates decides

Approximate solution for isotropic downsampling: Mipmaps

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering

Reconstruction is easy, prefiltering is difficult

Non-uniform resampling

- Reconstruct when locally upsampling
- Prefilter when locally downsampling
- Jacobian of map from screen to texture coordinates decides

Approximate solution for isotropic downsampling: *Mipmaps* 

Otherwise, use anisotropic filtering

# References

- E. C. Anderson. Monte carlo methods and importance sampling. UC Berkeley, 1999. Lecture notes for Stat 578C.
- T. Duff. Polygon scan conversion by exact convolution. In Jacques André and Roger D. Hersch, editors, *Raster Imaging and Digital Typography*, pages 154–168. Cambridge University Press, 1989.
- J. Manson and S. Schaefer. Analytic rasterization of curves with polynomial filters. *Computer Graphics Forum (Proceedings of Eurographics)*, 32(2pt4):499–507, 2013.
- D. Nehab and H. Hoppe. Random-access rendering of general vector graphics. ACM Transactions on Graphics (Proceedings of ACM SIGGRAPH 2008), 27(5):135, 2008.
- D. Nehab and H. Hoppe. A fresh look at generalized sampling.
  Foundations and Trends in Computer Graphics and Vision, 8(1):1–84, 2014.