## 2D Computer Graphics

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IMPA

## ANTI-ALIASING AND TEXTURE MAPPING

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How to compute the integral when $f$ is a vector graphics illustration?

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What about multiple opaque layers?
What about transparency?

## POPULAR HACK

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o=\int_{\Omega}\left[u-p \in P_{i}\right] \psi(u) d u
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The new background $b_{i+1}, \alpha_{i+1}$ is

$$
b_{i+1}, \alpha_{i+1}=f_{i},\left(\alpha_{i} \cdot 0\right) \oplus b_{i}, \alpha_{i}
$$

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Visible seams at perfectly abutting layers, weird halos


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Must blend in gamma and antialias in linear [Nehab and Hoppe, 2008]

$$
b_{i+1}, \beta i+1=\gamma\left(\gamma^{-1}\left(f_{i}, \alpha_{i} \oplus b_{i}, \beta_{i}\right) \cdot o+\gamma^{-1}\left(b_{i}, \beta_{i}\right) \cdot(1-0)\right)
$$

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i.e., the mean value weighted by the probability density function.

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The associated variance $\operatorname{var}(X)=\sigma_{X}^{2}$ is

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\operatorname{var}(X)=E\left[\left(X-\mu_{X}\right)^{2}\right]=E\left[X^{2}\right]-E^{2}[X]=\sigma_{X}^{2}
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Variance of sample average

$$
\operatorname{var}\left(\bar{X}_{n}\right)=\operatorname{var}\left(\frac{1}{n} \sum X_{i}\right)=\frac{1}{n^{2}} \sum \operatorname{var}\left(X_{i}\right)=\frac{\sigma_{x}^{2}}{n}
$$

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Let $X$ be such that support of $f_{X}$ is $\Omega$

$$
\int_{\Omega} g(t) d t=\int_{\Omega} \frac{g(t)}{f_{X}(t)} f_{X}(t) d t=E\left[\frac{g(X)}{f_{X}(X)}\right] \approx \frac{1}{n} \sum_{i=1}^{n} \frac{g\left(X_{i}\right)}{f_{X}\left(X_{i}\right)}
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This is the basis of supersampling
The solution to our anti-aliasing problems

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When $\psi=\beta^{0}$ is the box, $f_{X}=1$ with support $\Omega=\left[-\frac{1}{2}, \frac{1}{2}\right]^{2}$

$$
c(p) \approx \frac{1}{n} \sum_{i=1}^{n} g\left(p-X_{i}\right)
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The Monte Carlo estimator is unbiased in this sense

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It often makes sense to use a biased estimator to reduce variance

$$
c(p) \approx \frac{\sum_{i=1}^{n} \frac{g\left(p-X_{i}\right) \psi\left(X_{i}\right)}{f_{X}\left(X_{i}\right)}}{\sum_{i=1}^{n} \frac{\psi\left(X_{i}\right)}{f_{X}\left(X_{i}\right)}}
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This integral is exactly what we are trying to compute!

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This is importance sampling

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Many different point distributions have $f_{X}=1 / A_{\Omega}$ in $\Omega$
Uniform, stratified, low-discrepancy (e.g. Poisson disk, Lloyd relaxation) Variance of $\bar{X}_{n}$ is not the same for all of them!

16 SAMPLES

Regular


16 SAMPLES

Uniform

16 SAMPLES

Stratified


16 SAMPLES

Blue noise


## 64 SAMPLES

Regular


## 64 SAMPLES

Uniform

## 64 SAMPLES

Stratified

## 64 SAMPLES

Blue noise


Regular



## 256 SAMPLES

Stratified

Blue noise


## 1024 SAMPLES

Regular


## 1024 SAMPLES

## Uniform



## 1024 SAMPLES

## Stratified



## 1024 SAMPLES

Blue noise


## BETTER ANTI-ALIASING KERNELS





## BETTER ANTI-ALIASING KERNELS




Linear


## BETTER ANTI-ALIASING KERNELS




Gaussian



## BETTER ANTI-ALIASING KERNELS



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Cardinal B-spline





## Generalized sampling

discretization
reconstruction

$$
f_{\psi}=f * \psi^{v} \quad[\cdot] \quad \boldsymbol{c}=\left[f_{\psi}\right] * \boldsymbol{q} \quad \tilde{f}=\boldsymbol{c} * \varphi
$$



> continuous analysis

$\underset{\text { diltering }}{\text { digital }}$

mixed
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$$
\begin{aligned}
& \text { continuous } \\
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sampling
digital filtering

mixed synthesis

Cardinal cubic B-spline
Needs sample sharing for variance reduction and speed

## TEXTURING

Assuming good reconstruction and prefilter kernels,

- Upsampling needs only reconstruction
- Downsampling needs only prefiltering


## BOX UPSAMPLING



LINEAR UPSAMPLING


CARDINAL CUBIC B-SPLINE UPSAMPLING


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Approximate solution for isotropic downsampling: Mipmaps

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Approximate solution for isotropic downsampling: Mipmaps Otherwise, use anisotropic filtering

## References

E. C. Anderson. Monte carlo methods and importance sampling. UC Berkeley, 1999. Lecture notes for Stat 578C.
T. Duff. Polygon scan conversion by exact convolution. In Jacques André and Roger D. Hersch, editors, Raster Imaging and Digital Typography, pages 154-168. Cambridge University Press, 1989.
J. Manson and S. Schaefer. Analytic rasterization of curves with polynomial filters. Computer Graphics Forum (Proceedings of Eurographics), 32(2pt4):499-507, 2013.
D. Nehab and H. Hoppe. Random-access rendering of general vector graphics. ACM Transactions on Graphics (Proceedings of ACM SIGGRAPH 2008), 27(5):135, 2008.
D. Nehab and H. Hoppe. A fresh look at generalized sampling. Foundations and Trends in Computer Graphics and Vision, 8(1):1-84, 2014.

